Preferences erosion and the developing countries exports to the EU: a dynamic panel gravity approach

Valentina Raimondi (University of Milano)

Margherita Scoppola (University of Macerata)

Alessandro Olper (University of Milano)



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Motivation

- Trade liberalization may imply preferences erosion
- An interesting example is the EU rice policy: sharp reduction of MFN tariffs since 2004, preferential tariffs have been almost the same as before→ preferences erosion

The key policy issues:

 What is the size and what the trade impact of the preferences erosion occurred after the 2004 EU policy change?

Two contributions to the literature

 A new empirical approach to determine the preferential margin under tariff rate quotas which is consistent with economies of scale and imperfect competition

 A dynamic panel gravity model to account for endogeneity and persistency of trade with a quantitative variable (not a dummy) measuring preferences

The preferential margin (PM)

The PM of the *j* partner for the product *k* is:

$$PM_{kj} = \frac{T_k^{MFN} - T_{kj}^{PREF}}{1 + T_{kj}^{PREF}}$$

• T_{kj}^{PREF} : when preferences are granted by means of tariff rate quotas (TRQs) (i.e. two tariff system, with a tariff lower than the MFN one applied to the in-quota imports) What is their tariff equivalent?

The tariff equivalent of TRQs

To date the literature has determined the tariff equivalent of the TRQ on the basis of the perfect competition-increasing marginal cost model (marginal protection); the tariff equivalent is:

- the in-quota tariff if imports are lower than the quota;
- the out-of-quota tariff if imports are higher than the quota

However:

 Fixed costs and economies of scale prevail in the international trade of agricultural products

What is the tariff equivalent of a TRQ under economies of scale?

- We use the traditional Dixit-Stiglitz-Krugman model: symmetric firms importing a differentiated product, fixed costs + constant marginal cost → decreasing average costs; monopolistic competition with free entry ad exit;
- in equilibrium the price is equal to the average cost

The average cost under a TRQ is:

$$AC_{T^{in},T^{out}} = \begin{cases} \frac{FC}{Q} + c + \frac{T^{in}\overline{Q} + T^{out}(Q - \overline{Q})}{Q} & \text{if } Q > \overline{Q} \\ \frac{FC}{Q} + c + T^{in} & \text{if } Q \leq \overline{Q} \end{cases}$$

$$\emph{FC}$$
 = fixed cost \emph{c} = marginal cost; \emph{Q} = quota \emph{T}^{in} = in quota tariff \emph{T}^{out} = out-of-quota tariff, \emph{Q} : total imports

The tariff equivalent is:

 The weighted average of the two tariffs if imports are higher than the quota:

$$\frac{T^{in}\overline{Q} + T^{out}(Q - \overline{Q})}{Q} \text{ if } Q > \overline{Q}$$

• The in-quota tariff if imports are lower than the quota: T^{in} if $Q < \overline{Q}$

Hence, the tariff equivalent consistent with perfect competition if imports exceed the quota is higher than the one determined under economies of scale and monopolistic competition

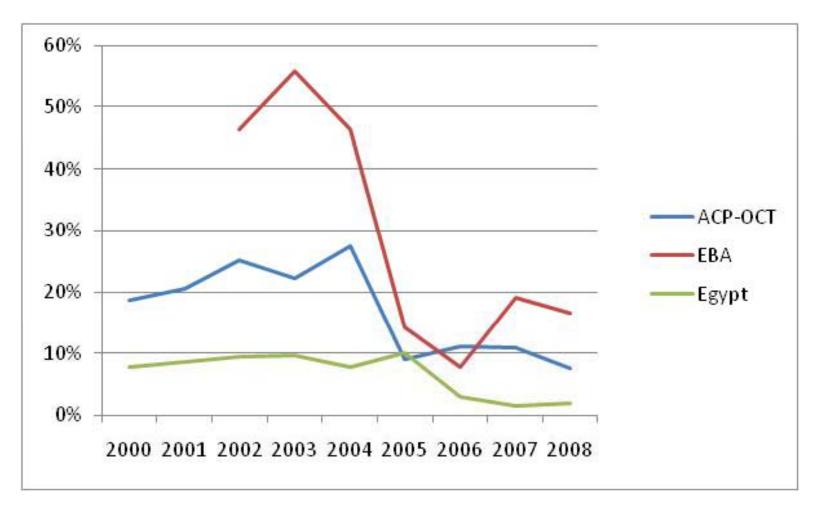
Preferential margins (PM) under the two hypotheses have been computed and then used into a gravity equation:

- PM^E under economies of scale
- PMP under perfect competition

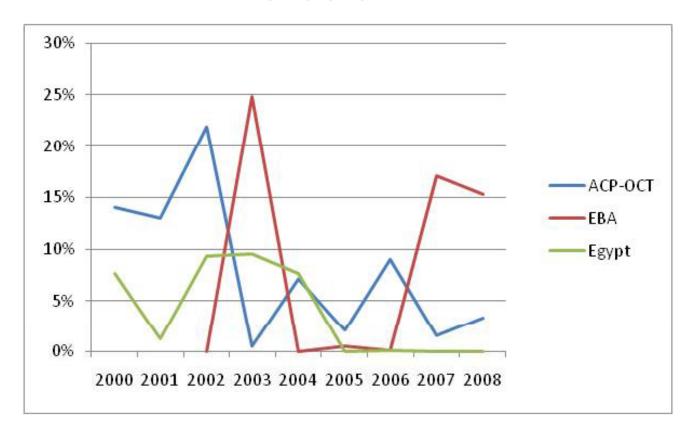
Data

- 9 years (2000-08) and 34 rice products (HS8 digit)
- 123 partner countries of the EU;
- In-quota and MFN tariffs converted in ad valorem tariffs (no aggregation);
- Total imports (HS8 digit) from Comext;
- In-quota imports from EC Commission

PM^E: evidence of preferences erosion after 2004 for EBA ad ACP countries



PM^P: NO clear evidence of preferences erosion!!



Assumptions about market structure and costs matter!

The empirical model

- Unobservable factors, often correlated with the level of trade, determine the governments' choice to select into a FTA
- This rises endogeneity in RHS variables
- Recent literature has shown that unbiased FTA effect can be estimated from (theoretically based) gravity models using panel data (e.g. Baier and Bergstrand, 2007)

The panel gravity model specification

$$\ln m_{ijkt} = \beta_0 + \beta_1 \ln(1 + T_{ijkt}) + \alpha_{jt} + \alpha_{it} + \alpha_{ij} + \alpha_{hs6t} + \alpha_t$$

where

- • m_{ijt} is the trade flow to country i from country j of good k in year t
- • T_{ijKt} is the ad valorem equivalent tariff
- • α_{it} and α_{jt} are the importer-year and exporter-year fixed effects
- • α_{ij} are bilateral fixed effects to control for unobserved time-invariant heterogeneities
- • α_t and α_{hs6t} are year and product-time dummies

The static panel gravity specification

$$\ln m_{ijkt} = \beta_0 + \beta_1 \ln \left(1 + T_{ijkt}\right) + \alpha_{jt} + \alpha_{it} + \alpha_{it} + \alpha_{hs6t} + \alpha_t$$

$$(1 + PM_{kj}) = \frac{(1 + T_k^{MFN})}{(1 + T_{kj}^{PREF})} \qquad \qquad (1 + T_{kjt}^{PREF}) = \frac{(1 + T_{kt}^{MFN})}{(1 + PM_{kjt})}$$

$$\ln m_{jkt} = \beta_0 + \beta_1 \left[\ln \left(1 + T_{kt}^{MFN} \right) - \ln \left(1 + PM_{jkt} \right) \right] + \alpha_{jt} + \alpha_{hs6t} + \alpha_t$$

$$\ln m_{jkt} = \beta_0 + \beta_2 \ln \left(1 + PM_{jkt}\right) + \alpha_{jt} + \alpha_{hs6t} + \alpha_t$$

Estimation Methods: LSDV for m > 0; Heckman; Poisson Pseudo Maximum Likelihood (PPML)

Dynamic gravity equation specification

To account for persistency, the equation is specified dynamically by adding the lagged dependent variable on the RHS.

$$\ln m_{jkt} = \beta_0 + \beta_1 \ln m_{jk(t-1)} + \beta_2 \ln (1 + PM_{jkt}) + \alpha_{jt} + \alpha_{hs6t} + \alpha_t + u_{jkt}$$

Moreover, given the structure of our panel (short-time series and large cross-section) we used the GMM estimator.

In particular, the System-GMM because of

- short panel data, along the time dimension (9 years)
- highly persistent data (trade flows)

Static model results - Panel (LSDV, Heckman, PPML)

| | D | ep. Variabl | Dep.Var.: Import _{jkt} | | | | |
|---------------------------|------------------------------|------------------------------|---------------------------------|------------------------------|------------------------------|------------------------------|--|
| _ | LS | LSDV | | HECKMAN | | PPML | |
| | Standard- PM ^P | Weighted- PM ^E | Standard- PM ^P | Weighted- PM ^E | Standard- PM ^P | Weighted- PM ^E | |
| | (1) | (2) | (3) | (4) | (5) | (6) | |
| log(1+PM _{jkt}) | 4.91** | 11.45*** | 20.54*** | 20.75*** | 10.64*** | 18.36*** | |
| i | (2.28) | (2.14) | (4.41) | (5.19) | (1.90) | (1.38) | |
| Mills ratio | | | 3.37*** | 1.85** | | | |
| | | | (0.74) | (0.82) | | | |
| No. of obs. | 3,195 | 3,195 | 3,195 | 3,195 | 17,944 | 17,944 | |

- The estimated preferences effect always increases in magnitude when the PM is measured assuming scale economies and imperfect competition (PM^E) (vis-à-vis perfect competition PM^E)
- Elasticity^{Heckman} > Elasticity^{PPML} > Elasticity^{LSDV}

Dynamic model results

(System-GMM)

| | Sys-GMM | | | |
|--|------------------------------|------------------|---|--|
| | Standard- PM ^P | Weighted PM | _ | |
| log(trade _{jk(t-1)}) | 0.61*** (0.10) | 0.61*** | | |
| log(1+PM _{jkt}) | 7.97 (7.64) | 5.03** (2.13) | | |
| log(distance _j) | -0.23 (1.96) | -0.17 (1.01) | | |
| log(production _{jt}) | 0.15 (0.11) | 0.16** (0.06) | | |
| No. Obs. AR(2) | 1,683 0.273 | 1,683 0.264 | | |
| Hansen p-value diff-in-Hansen p-value | 0.764 0.436 | 0.709 0.692 | | |

SHORT-RUN ELASTICITY

 A one percentage point increase in preferential margin factor is associated with a 5% increase in rice exports to the European Union, ceteris paribus.

LONG-RUN ELASTICITY

$$(\beta 2/(1-\beta 1))$$

- The long-run effect of the preferential margin factor on trade is near to 13
- The magnitude confirms the inertial behaviour of exports ₁₇

Dynamic model results in preferential groups

| | System- GMM | | | | | | | |
|--------------------------------|-------------|----------------------------|---------|---------|----------------------------|---------|--|--|
| | PN | PM ^P - Standard | | | PM ^E - Weighted | | | |
| | ACP-OCT | EBA | EGYPT | ACP-OCT | EBA | EGYPT | | |
| log(trade _{jk(t-1)}) | 0.70*** | 0.67*** | 0.65*** | 0.63*** | 0.67*** | 0.64*** | | |
| log(1+PM _{jkt}) | 2.40 | 3.74 | 17.27 | 10.36* | 3.70** | -3.82 | | |
| log(distance _j) | 0.13 | 0.16 | 0.16 | 0.11 | 0.13 | 0.07 | | |
| log(production _{it}) | 0.11*** | 0.13*** | 0.14*** | 0.16*** | 0.14*** | 0.14*** | | |

The **preferences impacts** estimated using PMP are <u>not significant</u>

PME * ACP-OCT countries have the <u>stronger impact</u> of preferences on trade

- * **EBA** countries have a <u>lower short-run</u> average effect of preferences on trade flows (preferences have drastically decreased after 2004)
- * not significant for **EGYPT** (increased ability to export broken-rice out of quota occurred after 2004)

Conclusions (1)

- The use of the <u>"standard" tariff equivalent</u> of tariff rate quotas may result in <u>misleading conclusions</u> about the extent of the trade preferences and of their impact on trade by:
 - an overestimation of the tariff equivalents and an underestimation of preferences
 - Empirically, this translates in an underestimation of trade elasticities to preferences (static model) or in not significant elasticities (dynamic model)
- Thus the assumptions about market structure and costs matter considerably when assessing the trade impact of preferences.

Conclusions (2)

From a policy perspective:

- Preferences erosion has been considerable according to the preferential margin consistent with economies of scale and monopolistic competition
- EU preferences still matter significantly on the developing countries ability to export rice to the EU
- and this is especially true for ACP countries

Thank you !!!!